

# The LASSO

CSci 8980: ML at Large Scale and High Dimensions

Instructor: Arindam Banerjee

January 29, 2014

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i)$ ,  $i = 1, \dots, n$ ,  $\mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i)$ ,  $i = 1, \dots, n$ ,  $\mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects
- Shrinking coefficients



# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects
- Shrinking coefficients
  - Increases bias, lowers variance, improves accuracy

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects
- Shrinking coefficients
  - Increases bias, lowers variance, improves accuracy
- Alternatives

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects
- Shrinking coefficients
  - Increases bias, lowers variance, improves accuracy
- Alternatives
  - Subset selection: Unstable, sensitive to small changes

# Regression with OLS

- Given training data  $(y_i, \mathbf{x}_i), i = 1, \dots, n, \mathbf{x}_i \in \mathbb{R}^p$
- Ordinary least squares (OLS)

$$\hat{\beta} = \operatorname{argmin}_{\beta \in \mathbb{R}^p} \sum_{i=1}^n (y_i - \beta^T \mathbf{x}_i)^2$$

- Issues/challenges with OLS
  - Accuracy: low bias, high variance
  - Interpretation: All coefficients are non-zero
  - Cannot determine small subsets with strong effects
- Shrinking coefficients
  - Increases bias, lowers variance, improves accuracy
- Alternatives
  - Subset selection: Unstable, sensitive to small changes
  - Ridge regression: Shrinks coefficients, but not to 0

# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0_j|$

# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0|$
- The non-negative garotte estimator (Breiman, 1996)

$$(\hat{\alpha}, \hat{c}) = \underset{(\alpha, c)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j c_j \hat{\beta}_j^0 x_{ij})^2 \text{ s.t. } c_j \geq 0, \sum_j c_j \leq t$$

# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0|$
- The non-negative garotte estimator (Breiman, 1996)

$$(\hat{\alpha}, \hat{c}) = \underset{(\alpha, c)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j c_j \hat{\beta}_j^0 x_{ij})^2 \text{ s.t. } c_j \geq 0, \sum_j c_j \leq t$$

- Relies on OLS  $\hat{\beta}^0$ : may be problematic in certain settings

# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0|$
- The non-negative garotte estimator (Breiman, 1996)

$$(\hat{\alpha}, \hat{c}) = \underset{(\alpha, c)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j c_j \hat{\beta}_j^0 x_{ij})^2 \text{ s.t. } c_j \geq 0, \sum_j c_j \leq t$$

- Relies on OLS  $\hat{\beta}^0$ : may be problematic in certain settings
- Least absolute shrinkage and selection operator (LASSO)

$$(\hat{\alpha}, \hat{\beta}) = \underset{(\alpha, \beta)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 \text{ s.t. } \sum_j |\beta_j| \leq t$$



# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0|$
- The non-negative garotte estimator (Breiman, 1996)

$$(\hat{\alpha}, \hat{c}) = \underset{(\alpha, c)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j c_j \hat{\beta}_j^0 x_{ij})^2 \text{ s.t. } c_j \geq 0, \sum_j c_j \leq t$$

- Relies on OLS  $\hat{\beta}^0$ : may be problematic in certain settings
- Least absolute shrinkage and selection operator (LASSO)

$$(\hat{\alpha}, \hat{\beta}) = \underset{(\alpha, \beta)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 \text{ s.t. } \sum_j |\beta_j| \leq t$$

- Parameter  $t < t_0$  will cause shrinkage

# The LASSO

- Let  $\hat{\beta}^0$  be the OLS solution, and  $t_0 = \sum_{j=1}^p |\hat{\beta}^0|$
- The non-negative garotte estimator (Breiman, 1996)

$$(\hat{\alpha}, \hat{c}) = \underset{(\alpha, c)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j c_j \hat{\beta}_j^0 x_{ij})^2 \text{ s.t. } c_j \geq 0, \sum_j c_j \leq t$$

- Relies on OLS  $\hat{\beta}^0$ : may be problematic in certain settings
- Least absolute shrinkage and selection operator (LASSO)

$$(\hat{\alpha}, \hat{\beta}) = \underset{(\alpha, \beta)}{\operatorname{argmin}} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 \text{ s.t. } \sum_j |\beta_j| \leq t$$

- Parameter  $t < t_0$  will cause shrinkage
  - Some coefficients will become 0

# Orthonormal Design Case

- Design matrix  $X \in \mathbb{R}^{n \times p}$ , assume  $X^T X = I \in \mathbb{R}^{p \times p}$

# Orthonormal Design Case

- Design matrix  $X \in \mathbb{R}^{n \times p}$ , assume  $X^T X = I \in \mathbb{R}^{p \times p}$
- Best subset selection picks  $k$  largest coefficients

# Orthonormal Design Case

- Design matrix  $X \in \mathbb{R}^{n \times p}$ , assume  $X^T X = I \in \mathbb{R}^{p \times p}$
- Best subset selection picks  $k$  largest coefficients
- For a suitable constant  $\gamma$ , the LASSO solution is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

# Orthonormal Design Case

- Design matrix  $X \in \mathbb{R}^{n \times p}$ , assume  $X^T X = I \in \mathbb{R}^{p \times p}$
- Best subset selection picks  $k$  largest coefficients
- For a suitable constant  $\gamma$ , the LASSO solution is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Ridge regression shrinks the coefficients

$$\hat{\beta}_j^{\text{ridge}} = \frac{1}{1 + \gamma} \hat{\beta}_j^0$$

# Orthonormal Design Case

- Design matrix  $X \in \mathbb{R}^{n \times p}$ , assume  $X^T X = I \in \mathbb{R}^{p \times p}$
- Best subset selection picks  $k$  largest coefficients
- For a suitable constant  $\gamma$ , the LASSO solution is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

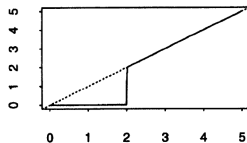
- Ridge regression shrinks the coefficients

$$\hat{\beta}_j^{\text{ridge}} = \frac{1}{1 + \gamma} \hat{\beta}_j^0$$

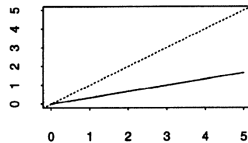
- Garotte estimates

$$\hat{\beta}_j^{\text{garotte}} = \left(1 - \frac{\gamma}{(\hat{\beta}_j^0)^2}\right)_+ \hat{\beta}_j^0$$

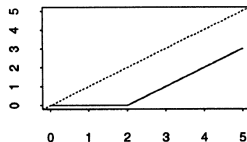
# Orthonormal Design Case



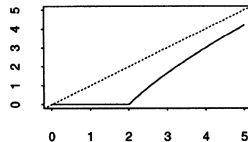
beta  
(a)



beta  
(b)



beta  
(c)



beta  
(d)

Shrinkage due to (a) subset selection, (b) ridge regression, (c) the lasso, and (d) the garrotte



- Elliptical contour of the objective

$$(\beta - \hat{\beta}^0)^T X^T X (\beta - \hat{\beta}^0)$$

- Elliptical contour of the objective

$$(\beta - \hat{\beta}^0)^T X^T X (\beta - \hat{\beta}^0)$$

- Level sets of the contour intersects with  $L_q$  norm ball

- Elliptical contour of the objective

$$(\beta - \hat{\beta}^0)^T X^T X (\beta - \hat{\beta}^0)$$

- Level sets of the contour intersects with  $L_q$  norm ball
  - $q = 2$ : Ridge regression, shrinkage but no sparsity

- Elliptical contour of the objective

$$(\beta - \hat{\beta}^0)^T X^T X (\beta - \hat{\beta}^0)$$

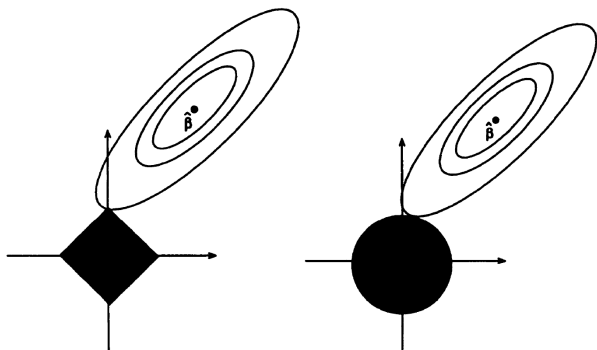
- Level sets of the contour intersects with  $L_q$  norm ball
  - $q = 2$ : Ridge regression, shrinkage but no sparsity
  - $q = 1$ : Lasso, shrinkage and sparsity

- Elliptical contour of the objective

$$(\beta - \hat{\beta}^0)^T X^T X (\beta - \hat{\beta}^0)$$

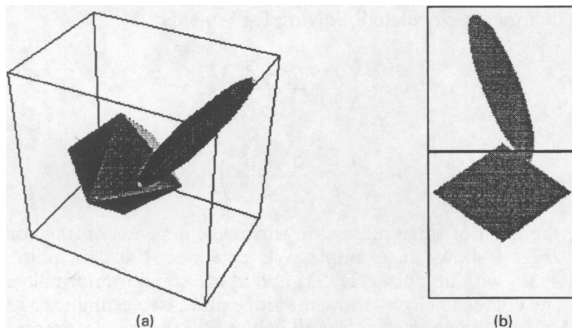
- Level sets of the contour intersects with  $L_q$  norm ball
  - $q = 2$ : Ridge regression, shrinkage but no sparsity
  - $q = 1$ : Lasso, shrinkage and sparsity
- Ridge vs Lasso: Can the sign change from OLS estimate?

# Geometry of LASSO: $p = 2$



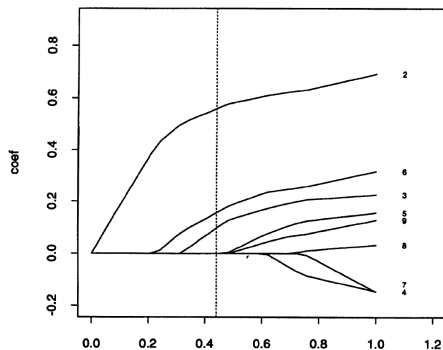
Estimation in (a) the lasso, and (b) ridge regression

# Geometry of LASSO: $p > 2$



Sign change in LASSO vs OLS is possible for  $p > 2$

# Example: Regularization Path



Shrinkage of parameters over  $s = \frac{t}{\sum_j \hat{\beta}_j^0}$



- The ‘regularized’ version of Lasso

$$(\hat{\alpha}, \hat{\beta}) = \operatorname{argmin}_{(\alpha, \beta)} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 + \lambda \sum_j |\beta_j|$$

# Estimating “t”

- The ‘regularized’ version of Lasso

$$(\hat{\alpha}, \hat{\beta}) = \operatorname{argmin}_{(\alpha, \beta)} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 + \lambda \sum_j |\beta_j|$$

- Cross-validation over  $\lambda$  (or  $t$ )

- The ‘regularized’ version of Lasso

$$(\hat{\alpha}, \hat{\beta}) = \operatorname{argmin}_{(\alpha, \beta)} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 + \lambda \sum_j |\beta_j|$$

- Cross-validation over  $\lambda$  (or  $t$ )
  - Pick the value that leads to smallest error

# Estimating “ $t$ ”

- The ‘regularized’ version of Lasso

$$(\hat{\alpha}, \hat{\beta}) = \operatorname{argmin}_{(\alpha, \beta)} \sum_{i=1}^n (y_i - \alpha - \sum_j \beta_j x_{ij})^2 + \lambda \sum_j |\beta_j|$$

- Cross-validation over  $\lambda$  (or  $t$ )
  - Pick the value that leads to smallest error
- Resampling based estimates, e.g., stability selection

# Generalized Regression Models

- General regression problem formulation

# Generalized Regression Models

- General regression problem formulation
  - Constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) \quad \text{s.t.} \quad \|\beta\|_1 \leq t$$

# Generalized Regression Models

- General regression problem formulation
  - Constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) \quad \text{s.t.} \quad \|\beta\|_1 \leq t$$

- Regularized version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) + \lambda \|\beta\|_1$$

# Generalized Regression Models

- General regression problem formulation

- Constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) \quad \text{s.t.} \quad \|\beta\|_1 \leq t$$

- Regularized version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) + \lambda \|\beta\|_1$$

- The other constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} \|\beta\|_1 \quad \text{s.t.} \quad L(y, X, \beta) \leq a$$



# Generalized Regression Models

- General regression problem formulation

- Constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) \quad \text{s.t.} \quad \|\beta\|_1 \leq t$$

- Regularized version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) + \lambda \|\beta\|_1$$

- The other constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} \|\beta\|_1 \quad \text{s.t.} \quad L(y, X, \beta) \leq a$$

- Examples: logistic regression, generalized linear models, etc.

# Generalized Regression Models

- General regression problem formulation

- Constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) \quad \text{s.t.} \quad \|\beta\|_1 \leq t$$

- Regularized version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} L(y, X, \beta) + \lambda \|\beta\|_1$$

- The other constrained version

$$\hat{\beta} = \underset{\beta}{\operatorname{argmin}} \|\beta\|_1 \quad \text{s.t.} \quad L(y, X, \beta) \leq a$$

- Examples: logistic regression, generalized linear models, etc.
- We will consider efficient algorithms for such general problems

- Consider orthonormal design  $X^T X = I$ , so Lasso estimate is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Consider orthonormal design  $X^T X = I$ , so Lasso estimate is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Let  $\beta$  be the 'true' parameter:

$$y = \beta^T \mathbf{x} + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- Consider orthonormal design  $X^T X = I$ , so Lasso estimate is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Let  $\beta$  be the 'true' parameter:

$$y = \beta^T \mathbf{x} + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- Consider risk

$$R(\hat{\beta}, \beta) = E\|\hat{\beta} - \beta\|^2$$

- Consider orthonormal design  $X^T X = I$ , so Lasso estimate is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Let  $\beta$  be the 'true' parameter:

$$y = \beta^T \mathbf{x} + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- Consider risk

$$R(\hat{\beta}, \beta) = E\|\hat{\beta} - \beta\|^2$$

- Let  $R_{DP}$  be the loss of the 'optimal' predictor

$$T_{DP}(\hat{\beta}^0, \delta) = (\delta_j \hat{\beta}_j^0), \quad \delta_j = I(|\beta_j| > \sigma) \in \{0, 1\}$$

- Consider orthonormal design  $X^T X = I$ , so Lasso estimate is

$$\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$$

- Let  $\beta$  be the 'true' parameter:

$$y = \beta^T \mathbf{x} + \epsilon, \quad \epsilon \sim N(0, \sigma^2)$$

- Consider risk

$$R(\hat{\beta}, \beta) = E\|\hat{\beta} - \beta\|^2$$

- Let  $R_{DP}$  be the loss of the 'optimal' predictor

$$T_{DP}(\hat{\beta}^0, \delta) = (\delta_j \hat{\beta}_j^0), \quad \delta_j = I(|\beta_j| > \sigma) \in \{0, 1\}$$

- $T_{DP}$  needs knowledge of  $\beta$ , not practical

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$



# Bounds on the Risk: Donoho et al.

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk
- Soft threshold estimator  $\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk
- Soft threshold estimator  $\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$ 
  - With  $\gamma = \sigma(2 \log n)^{1/2}$ , has same behavior

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk
- Soft threshold estimator  $\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$ 
  - With  $\gamma = \sigma(2 \log n)^{1/2}$ , has same behavior
- General design matrices

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk
- Soft threshold estimator  $\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$ 
  - With  $\gamma = \sigma(2 \log n)^{1/2}$ , has same behavior
- General design matrices
  - Lasso estimator continues to have good properties

- Hard threshold estimator  $\tilde{\beta}_j = \hat{\beta}_j^0 I(|\hat{\beta}_j^0| > \gamma)$ 
  - Has risk  $R(\tilde{\beta}, \beta) \leq (2 \log p + 1)(\sigma^2 + R_{DP})$
  - Threshold  $\gamma = \sigma(2 \log n)^{1/2}$  to get smallest asymptotic risk
- Soft threshold estimator  $\hat{\beta}_j = \text{sign}(\hat{\beta}_j^0)(|\hat{\beta}_j^0| - \gamma)_+$ 
  - With  $\gamma = \sigma(2 \log n)^{1/2}$ , has same behavior
- General design matrices
  - Lasso estimator continues to have good properties
  - Generalized to other sparsity inducing norms

# Norm level sets



(a)



(b)



(c)



(d)



(e)

$L_q$  norm level sets: (a)  $q = 4$ , (b)  $q = 2$ , (c)  $q = 1$ , (d)  $q = 0.5$ ,  
(e)  $q = 0.1$